

# Spatial Statistics of the CMB

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# Outline

- 1 Introduction
- 2 Symmetries
- 3 Interpretation

# Random Fields

## Definition

A random field is a spatial field with an associated probability measure:  $\mathcal{P}(A)DA$ .

- Random fields are abundant in Cosmology.
- The cosmic microwave background fluctuations constitute a random field on a sphere.
- Other examples: Dark Matter Distribution, Galaxy Distribution, etc.
- Astronomers measure particular realization of a random field (ergodicity helps but we cannot avoid “cosmic errors”)

# Isotropy – Anisotropy

## Definition

The CMB statistics is assumed to be invariant under  $(SO(3))$  rotations of the sky.

- This follows from the cosmological principles of homogeneity and isotropy.
- CMB temperature fluctuations are called “anisotropy”
- The anisotropy is statistically isotropic.

# Gaussian – Non-Gaussian

## Higher Order Spatial Statistics

Studies and characterizes the spatial distribution random fields beyond Gaussianity.

- Gaussian: statistics up to second order.
- Non-Gaussian: connected moments are non-zero beyond second order.
- Gaussian/Non-Gaussian: correlated/random phases.
- The CMB is Gaussian to a high degree of accuracy
- Isotropic non-Gaussian or Anisotropic Gaussian?

# Definitions

- The ensemble average  $\langle A \rangle$  corresponds to a functional integral over the probability measure.
- Physical meaning: average over independent realizations.
- Ergodicity: (we hope) ensemble average can be replaced with spatial averaging.

## Joint Moments

$$F^{(N)}(x_1, \dots, x_N) = \langle T(x_1), \dots, T(x_N) \rangle$$

# Connected Moments

These are the most frequently used spatial statistics

- Typically we use fluctuation fields  $\delta = T/\langle T \rangle - 1$

Connected moments are defined recursively

$$\langle \delta_1, \dots, \delta_N \rangle_c = \langle \delta_1, \dots, \delta_N \rangle - \sum_P \langle \delta_1 \dots \delta_i \rangle_c \dots \langle \delta_j \dots \delta_k \rangle_c \dots$$

- With these the  $N$ -point correlation functions are

$$\xi^{(N)}(1, \dots, N) = \langle \delta_1, \dots, \delta_N \rangle_c$$

# Basic Objects

These are  $N$ -point correlation functions.

## Special Cases

Two-point functions	$\langle \delta_1 \delta_2 \rangle$
Three-point functions	$\langle \delta_1 \delta_2 \delta_3 \rangle$
Cumulants	$\langle \delta_R^N \rangle_c = \mathcal{S}_N \langle \delta_R^2 \rangle^{N-1}$
Cumulant Correlators	$\langle \delta_1^N \delta_2^M \rangle_c$
Conditional Cumulants	$\langle \delta(0) \delta_R^N \rangle_c$

- In the above  $\delta_R$  stands for the fluctuation field smoothed on scale  $R$  (different  $R$ 's could be used for each  $\delta$ 's).
- Host of alternative statistics exist: e.g. Minkowski functions, void probability, minimal spanning trees, phase correlations, etc.

# Complexities

Combinatorial explosion of terms

- $N$ -point quantities have a large configuration space: measurement, visualization, and interpretation become complex.
- CPU intensive measurement:  $M^N$  scaling for  $N$ -point statistics of  $M$  objects.
- Theoretical estimation
- Estimating reliable covariance matrices
- As we will see, a close look at symmetries helps with these issues.

# Symmetries in General

Invariance under a (Lie) group  $G$

Expansion according to irreducible representations

$$\begin{aligned}\xi(gx, gy) &= \xi(x, y) \\ \delta(x) &= \sum_i a_i^\alpha \psi_i^\alpha\end{aligned}$$

Irreducible reps diagonalize corr. matrix

$$\begin{aligned}\langle a_i^\alpha a_j^\beta \rangle &= \delta_{ij} \delta_{\alpha\beta} \int dx \xi(0, x) \sum_s \psi_s^\alpha(0) \psi_s^{\alpha*}(x) \frac{1}{f_\alpha} \int dg \\ \langle a_{l_1 m_1} a_{l_2 m_2} \rangle &= \delta_{l_1 l_2} \delta_{m_1 m_2} C_{l_1}\end{aligned}$$

# Expansions for the sphere

Statistics under  $SO(3)$  invariance

## Angular Power Spectrum

$$\xi(\theta) = \sum \frac{2l+1}{4\pi} C_l P_l(\cos \theta)$$

- Legendre polynomials  $P_l \propto$  bipolar spherical harmonics  
 $L = 0$

## Bispectrum: Tripolar Expansion

$$\begin{aligned} S_{l_1 l_2 l_3}(\hat{x}_1, \hat{x}_2, \hat{x}_3) &\equiv \sum_{m_1, m_2, m} \begin{pmatrix} l_1 & l_2 & l_3 \\ m_1 & m_2 & m_3 \end{pmatrix} Y_{l_1 m_1}(\hat{x}_1) Y_{l_2 m_2}(\hat{x}_2) Y_{l_3 m_3}(\hat{x}) \\ \xi^3(\hat{x}_1, \hat{x}_2, \hat{x}_3) &= \sum_{l_1, l_2, l_3} B_{l_1, l_2, l_3} S_{l_1 l_2 l_3}(\hat{x}_1, \hat{x}_2, \hat{x}_3) \end{aligned}$$

# Algorithms: SpICE

Szapudi Etal 2001a,b

- `www.ifa.hawaii.edu/cosmowave`
- about 258 downloads to date
- Heuristic weighting/pseudo- $C_l$  method
- Use fast harmonic transform (HEALPix Gorski etal 2005) to calculate raw two-point correlators: 2 minutes for  $N_{sdl} = 512$ .
- The coupling kernel is diagonal in pixel space: invert there
- Use Gauss-Legendre quadrature to calculate  $C_l$ 's
- Monte Carlo simulations for covariance matrices

## Algorithms: SpICE cont'd

- Equivalent technique (MASTER Hivon Etal 2002) directly inverts the (non-diagonal) coupling matrix
- reanalysed WMAP-I with SpICE (Fosalba & Szapudi 2004) coupling matrix in harmonic space
- Part of Planck level 2 pipeline
- Works for  $N - M$  cumulant correlators (Szapudi & Szalay 1992) as well
- Direct generalization to bispectrum is not fast enough ( $N^{5/2}$ ) and would calculate  $l^3 \simeq N^{3/2}$  bins

# Flat Sky Approximation: Translation Invariance

The Dirac  $\delta$ 's are expressing the symmetry

## Fourier Transform

$$\begin{aligned}\langle \delta(k_1)\delta(k_2) \rangle &= (2\pi)^D \delta_D(k_1 + k_2) P(k_1) \\ \langle \delta(k_1)\delta(k_2)\delta(k_3) \rangle &= (2\pi)^D \delta_D(k_1 + k_2 + k_3) B(k_1, k_2, k_3)\end{aligned}$$

## Generalized Wiener-Khinchin Theorem

$$\begin{aligned}\xi(x_1, x_2) &= \int \frac{d^D k}{(2\pi)^D} P(k) e^{ik(x_1 - x_2)} \\ \xi(x_1, x_2, x_3) &= \int \prod_{i=1}^3 \frac{d^D k_i}{(2\pi)^D} B(k_1, k_2, k_3) e^{i(\sum k_j x_j)} \delta_D(\sum k_j).\end{aligned}$$

# Symmetries: Rotation Invariance

These further simplify things...

- The two-point function depends on the distance of the two variables.

The integral is 1D only

$$\xi(r) = \int \frac{kdk}{2\pi} P(k) J_0(kr) \quad 2D,$$

$$\xi(r) = \int \frac{k^2 dk}{2\pi^2} P(k) j_0(kr) \quad 3D$$

- Can we use the same idea for the 3 point statistics?
- They depend on a triangle size and shape.

# Rotation Invariance: 3D

SO(3)

- If we parametrize the triangle by two unit vectors, and two sizes, bipolar spherical expansion can capture the rotation invariance.
- In the bipolar expansion only  $L = 0$  (scalar) can happen because of SO(3) invariance: Legendre polynomials.

## Multipole Expansion For 3-point Statistics

$$B(k_1, k_2, \theta) = \sum_l B_l(k_1, k_2) P_l(\cos \theta) \frac{2l + 1}{4\pi},$$

- Note that tripolar expansion is also possible and more convenient in some cases.

## Consequence of Rotation Invariance in 3D

- The above expansion simplifies the connection between Fourier and real space 3-point statistics.
- In particular the 6 dimensional (intractable integral) becomes 2 dimensional.

### Double Hankel Transform

$$\xi_l^3(r_1, r_2) = \int \frac{k_1^2}{2\pi^2} dk_1 \frac{k_2^2}{2\pi^2} dk_2 (-1)^l B_l(k_1, k_2) j_l(k_1 r_1) j_l(k_2 r_2)$$

- Multiples do not mix in the transformation!

# Rotational Invariance in 2D

## U(1) invariance

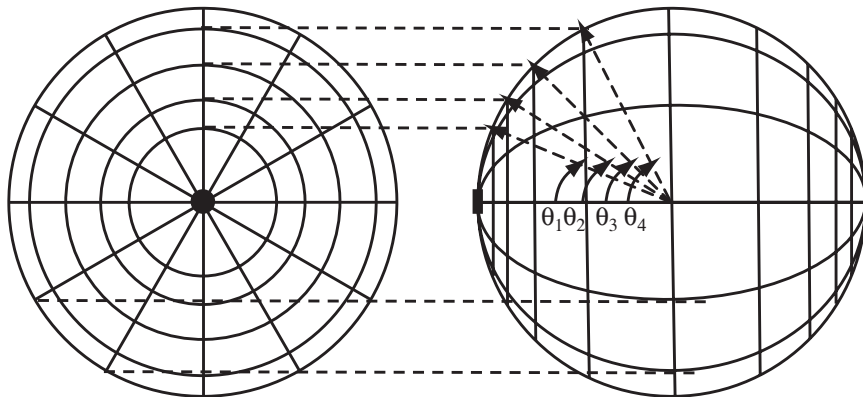
- In 2 dimensions  $SO(3) \rightarrow U(1)$ , we need to use cosine expansion.

### 2D Expansion

$$B(k_1, k_2, \theta) = \frac{B_0(k_1, k_2)}{2} + \sum_{n < 0} B_n(k_1, k_2) \cos(n\theta)$$
$$\xi_n^3(r_1, r_2) = \int \frac{k_1}{2\pi} dk_1 \frac{k_2}{2\pi} dk_2 (-1)^n B_n(k_1, k_2) J_n(k_1 r_1) J_n(k_2 r_2)$$

- The same expansion can be applied to spherical geometry in the flat sky limit, and for the (redshift space) projected three-point function.

# New Algorithm for 3pt

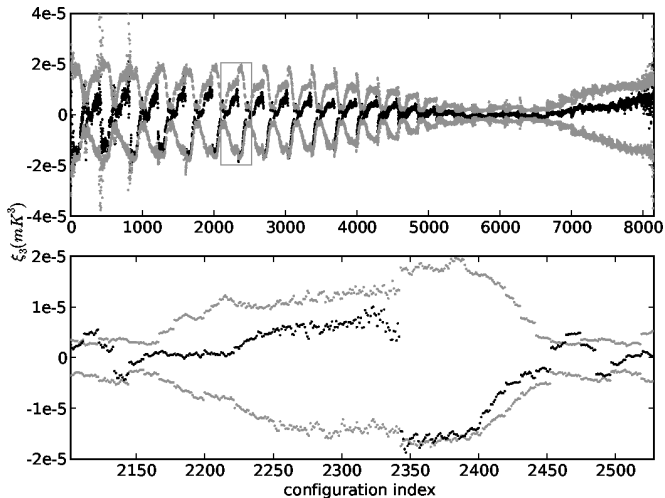


## New Algorithm for 3pt Cont'd

- Naively  $N^3$  calculations to find all triplets in the map: overwhelming (millions of CPU years for WMAP)
- Regrid CMB sky around each point according to the resolution
- Use hierarchical algorithm for regridding:  $N \log N$
- Correlate rings using FFT's (total speed: 2 minutes/cross-corr)
- The final scaling depends on resolution  

$$N(\log N + N_\theta N_\alpha \log N_\alpha + N_\alpha N_\theta (N_\theta + 1)/4)/2$$
- With another cos transform one and a double Hankel transform one can get the bispectrum
- In WMAP-I: 168 possible cross correlations, about 1.6 million bins altogether.
- How to interpret such massive measurements?

# 3pt in WMAP

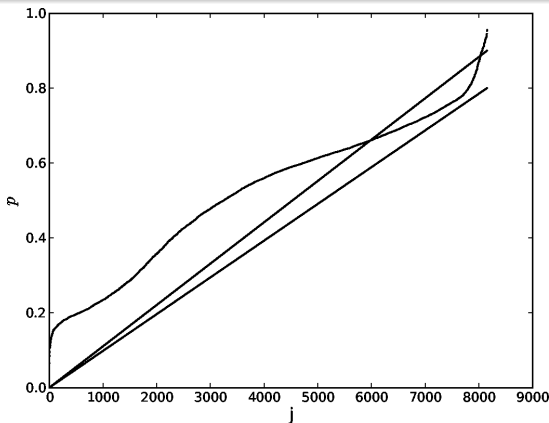


# Question 1: How consistent is the CMB with Gaussian?

## Hypothesis testing with FDR

- Gaussianity: null hypothesis
- Combining 1.6 million (correlated) measurements
- Benjamini and Hochberg (1995), Benjamini and Yekutieli (2001)
- Miller et al (2001) Chen and Szapudi (2005)
- False Discovery Rate: fix  $\alpha$  s.t.  $0 \leq \alpha \leq 1$
- $P_1, \dots, P_N$  sorted  $p$ -values
- let  $d = \max \left\{ j : P_j < \frac{j\alpha}{c_N N} \right\}$
- Configurations with  $i < d$  are rejected, and  $\langle FDR \rangle \leq \alpha$  is guaranteed
- $c_N \lesssim \sum_{i=1}^N i^{-1}$  for correlated data, otherwise 1 (Hopkins et al 2002)

# FDR in WMAP



- $\alpha = 0.8$  produces no rejection ( $\alpha = 0.9$  produces many rejections)
- the data are fully consistent with Gaussianity

# The Question: Are there Primordial Gaussianities?

## Parameter Fitting

- It is difficult to translate FDR results into traditional confidence levels
- Our goal is to constrain primordial non-Gaussianities (others: e.g., point sources, foregrounds, SZ, lensing, etc)
- Since three-point correlations are expected to dominate (unless some symmetry erases three-point correlations)
- $\langle \delta^3 \rangle \ll \langle \delta^2 \rangle$

### Simplest Phenomenological Parametrization

$$\delta = \delta_G + f_{NLT}(\delta_G^2 - \langle \delta_G^2 \rangle)$$

# A Popular Parametrization

Phenomenology motivated by inflationary models

- $\Phi(k)$  is the primordial curvature perturbation

## Perturbative Model

$$\Phi(x) = \Phi_L(x) + f_{NL}(\Phi(x)_L^2 - \langle \Phi_L^2(x) \rangle)$$

- The leading order effect will be bispectrum, (three-point correlations)

## Bispectrum (Komatsu and Spergel 2001)

$$b_{l_1 l_2 l_3} = 2 \int r^2 dr (b_{l_1}^L(r) b_{l_2}^L(r) b_{l_3}^{NL}(r) + \text{perm.})$$

- where  $b^L$  and  $b^{NL} = \frac{2}{\pi} \int k^2 dk f(k) g_{\mathcal{T}l}(k) j_l(k)$  with  $f = P_\Phi(k)$  or  $f_{NL}$  respectively.
- Simple calculation through modification of CAMB (or CMBFAST).

# Measurements of $C_l^{2,1}$ in WMAP-I and WMAP-II

- SpICE
- Flat weights (we have tried other weights and it does not make any difference)
- Monte Carlo simulations for covariance matrices
- Gaussian realizations with “Best Fit”  $C_l$ 's
- Conservative Kp2 mask
- 110 Noise maps for WMAP-I, 200 self generated noise maps for WMAP-II
- Results yield about 1000x168 bins for each simulation or 16800  $C_l^{2,1}$  measurements.

# Theoretical Predictions

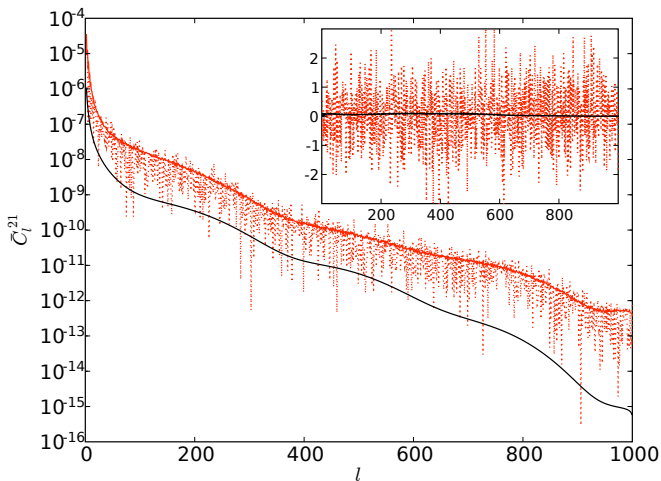
- Theory scales linearly with  $f_{NL}$ .
- Modified CAMB to obtain the bispectrum

Cooray (2001)

$$C_l^{2,1} = \sum_{l_1 l_2} B_{l_1 l_2 l} w_{l_1, l_2} \begin{pmatrix} l_1 & l_2 & l \\ 0 & 0 & 0 \end{pmatrix} \sqrt{\frac{(2l_1 + 1)(2l_2 + 1)}{4\pi(2l + 1)}}$$

- Predictions include beam and pixel window functions

# $C_l^{2,1}$ Measurements from WMAP-II



# Maximum Likelihood Fit

## Schematic Plan

- Estimate covariance matrix  $C$  of the measurements using a set of Monte Carlo simulations of the data
- Since non-Gaussianity is small, most noise is expected to come from the Gaussian variance
- Gaussian simulations are adequate if  $f_{NL}$  is small
- Estimate  $\chi^2 \simeq y^T C^{-1} y$ , where  $y = C_l^{2,1}(WMAP) - C_l^{2,1}(f_{NL})$
- Minimize  $\chi^2$  to find the best fit value and the variance of the parameter

## Problem 1: too many bins in three-point functions

- About 1.6 million bins for limited resolution, full three-point analysis.
- 168k bins for full resolution  $C_l^{2,1}$ , although it can be decreased by using  $l$ -bands and averages.
- Averaging cross-correlations and band powers can help that.
- This can be a problem for any statistics with many bins: e.g. full resolution cross- $C_l$ 's between many maps.

# Covariance Matrix Too Large

- Direct inversion of a  $10^6 \times 10^6$  matrix is not feasible, 40-50k matrix is doable, although inconvenient.
- It turns out not even needed!
- Let  $M = (p \times q)$  simulation matrix with  $p$  is the number of realizations,  $q$  is the number of bins (length of the result vector).
- $C = MM^T/p$
- Consider the SVD  $M = U\Lambda V$ , then  $MM^T = U\Lambda^2U^T$
- Solve the dual problem first  $M^TM = V^T\Lambda^2V$ , then find the eigen-vectors  $U$  from  $MV^T$ .
- In our case the problem is reduced to inverting a  $110 \times 110$  matrix.
- **Corollary** : even theoretically, only up to  $p$  modes can be extracted.

## Problem 2: Too Few Simulations

- At most 110 noise realizations for WMAP-I are available
- Hundreds (maybe thousands) of simulations are feasible, for WMAP-II but not millions or billions, even with fast codes.
- For convergence we need  $p \gg q$ : unfeasible.
- The problem is slightly less severe for  $C_l$ , but it is there: we are never in the  $p \gg q$  limit
- A model dependent remedy is to fit a simple analytical form

# Random Matrix Theory

Initiated by Jenő Wigner

- $M_{ia}M_{jb} = \sigma^2\delta_{ij}\delta_{ab}$

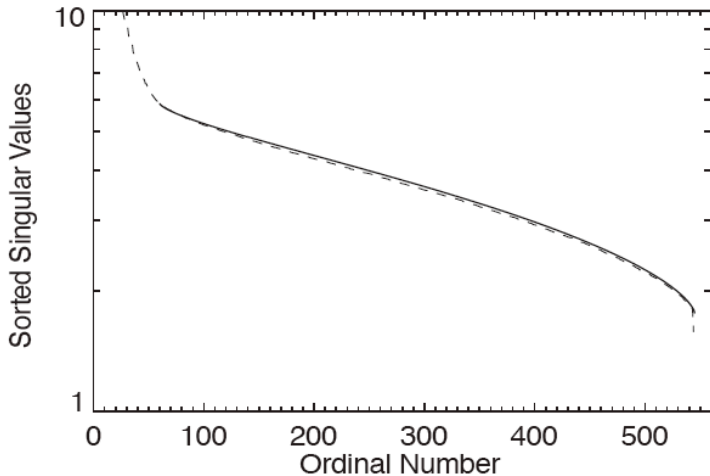
## Eigenvalue Spectrum

$$\rho(\lambda) = \frac{q^{1/2}}{\pi\lambda\sigma} \sqrt{(\lambda_{max}^2 - \lambda^2)(\lambda^2 - \lambda_{min}^2)}$$

$$\lambda_{max,min} = \sqrt{2}\sigma \sqrt{(p+q)/2 \pm \sqrt{pq}}$$

- Sengupta and Mitra (1997)

## Typical Situation

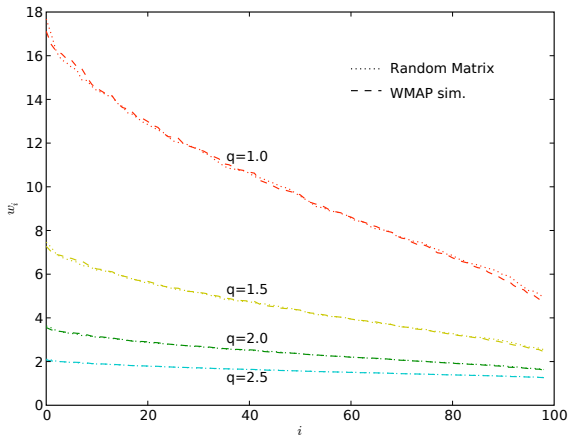


# Power Mapping

Guhr & Kälber (2003)

- A way to decrease the noise and identify significant eigenmodes
- $C_{kl} \rightarrow \text{sgn}(C_{kl}) | C_{kl} |^r$
- variance scales as  $1/p^{r/2}$
- e.g., 100 simulations with  $r = 2$  scatters like 10,000 simulations
- A series of power maps will reveal hidden structure in otherwise random looking covariance matrices.
- This is more an art than science...

# WMAP-II Covariance Matrix vs. Random



# The Zen of the Covariance Matrix

- The WMAP-II (and also WMAP-I) covariance matrices are consistent with random
- It is correct to take into account covariance in the  $\chi^2$  when those covariances are well known.
- It can be misleading to use the noisy off diagonal terms in a random matrix
- Investment banks paid for this experience dearly
- in this case it appears to be best to use a diagonal  $\chi^2$ , which is fully consistent with the randomness of the covariance
- One possible refinement is to fit  $N_{eff}$  d.o.f. from the shape of the curve.

# Constraints for $f_{NL}$ from $C_l^{2,1}$ in WMAP-II

These results are preliminary

- We use the diagonal elements of covariance matrix: the rest is consistent with noise.
- Using an average over the 168 cross-correlations we obtain  $f_{NL} = 22 \pm 52$  (68%)
- These results are comparable with the results in Spergel etal (2006)
- No optimal weighting, but cross correlations and more configuration dependence
- While we used accurate masks which take care of most of the point source contribution, we did not subtract possible contribution from unresolved point sources.

# Summary

## Three-point statistics of the CMB

- Most of the information in the CMB is in the power spectrum
- Did not talk about parameter estimation from the 2pt and polarization (both are hot topics though)
- Higher order statistics is especially difficult for CMB
- Careful look at the symmetries and the perturbative nature of possible non-Gaussianities help with estimators and algorithms
- The CMB is fully consistent with Gaussian, both in terms of hypothesis testing and parameter fitting